Economic Intuition to Social Capital: Household Evidence from Jimma Zone, South-West Ethiopia

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ABSTRACT

The notion of social capital is noble in social science literatures and recently getting noteworthy attention in economics. Its economic relevance has also been well documented in empirical studies. In this paper, we established relationships between the stock of household social capital and annual net income following approach for binary logit. Besides, we examined the relative economic importance of general vis special stocks of social capital. Statistical requirements with logistic regression were all tested robust. Consequently, the effect of social capital stock on household economy has found substantial. Of the two distinct forms, the special facet of social capital was found better and significant contributor to household economy. However, the general aspect was suggested insignificant as it was not meaningfully interpreted in the household economic network. Thus, a wise household need to appropriate the general to special social capital establishment, since every additional stock of the later could be meant to considerably help own welfare.

Keywords: Household Economy, Jimma Zone, Logistic analysis, Social Capital Stock

INTRODUCTION

Despite its multifaceted insinuation to nations' socioeconomic development, the topic of social capital has given trifling attention in economics. Without upheld social capital, it's hard to think of developments in the rest of the economy as institutional quality is a priori. A growing number of studies divulge an essential link between social capital measures and several indicators of economic outcomes. Wickramasuriya et al. (2011) well-thought-out social capital a moment they found significant role of civic engagement and enhanced personal networks in their household level analysis from Central Province of Sri Lanka. We need to be clear that, the idea of social capital, here, constitute personal, or either, engagements in civic organizations and established inter/and-intra-personal interconnections. Obviously, both could serve a promising part of societal networks in a meaningful manner. Coleman (1999) has defined ‘social capital’ as a multiplicity of various entities, having two mutual elements; all consisting of some feature of social buildings. According to him, these mutual elements ease certain activities of individuals belonging in the structure through connections of trust, interchange and interactions. Bowles and Gintis (2002) referred ‘trust, individual’s concern
one another, individual’s willingness to be governed by norms and cultures of general community, and one’s commitment to punish others who malfeasance the established norm”. Today, it is decisively rooted in economics and other social science disciplines. The term ‘social capital’ has lately linked more to human network and relational lives rather than social norms. For example (Stiglitz, 1999; Paldam, 2000; Dasgupta, 2005; Sobel, 2002) defined the term ‘social capital’ as a collection of social relations. Social networks undoubtedly entail several investment and even sacrifice in the form of money, time, prestige and information in a manner that could group access to better income, social life, knowledge and other benefits (Uphoff, 1999; Stiglitz, 1999; Dasgupta, 2005). Some networks impede free of cost; naturally you were born default within certain relations whether you like it or not. These types of networks normally are known as general relations. However, other forms of relationships can be established through a costly process, and they often have meaningful interpretation in life. In literature, the later forms are known as special relations (Bowles and Gintis, 2002; Dasgupta, 2005; Putnam, 1995; Ferragina, 2010).

An economic role of social capital is well documented in social science and economic literatures. In their national-wise survey Knack and Keefer (1997) has, in turn, shown that a unit measure scale advance in state level trust could result in a more than three-fold improvement of economic growth. The evidence shows actuality of large response from aggregate economy with respect to changes in national stock of social capital. In this paper we are not going through a more advanced model for social capital since it is limited to micro-level analysis. We rather rely on simple definition of social capital emphasizing norms, institutions and conventions. Avoiding complexities, we consider individuals’ civic engagement dummy and personal interrelationships to satisfy the definition of social capital due to (Miller, 2003; Greene and Brown, 1997; Becker, 1964; Mencer, 1974; Woolcock, 1998; Schuller et al., 2000; Coleman, 1988).

Empirical evidences on the relations between social capital and economic outcomes have clear theoretical foundation. Yasunobu and Bhandari (2009) demonstrate the positive role of social capital on economic growth since social ties intensify efficiency across beneficiary group. With enhanced cooperation and civic engagement, it is likely that both production and distribution costs could decrease. Since marketing information is now expected to be availed easily, transaction costs would also vanish. Moreover, effective team work is supposed to enhance innovation as well as exchanges of knowledge across units. All this effects from social capital is acknowledged to enhance productivity and growth at the same time (Fudenberg and Maskin, 1986; Yasunobu and Bhandari, 2009).

The Organization for Economic and Cultural Development (OECD) demonstrate the positive aspect of social capital on economic growth through potential efficiency that can be realized from societal cooperation (OECD, 2013). This part of demonstration bases on the fact that, the networks of norms swift trust and ensure collaboration. Theoretically sound is to assume operational efficiency, and hence growth, that could result from enhanced networks and cooperation. This paper also lays its analysis on this theoretical foundation regarding the topic we are dealing with.
We survey a random draw of households from various formations of Jimma zonal state administration. It, in the first place, is initiated to lesson whether or not household stock of social capital has to do with own economic status. Thus, our survey should be understood as micro-unit analysis unlike to most literatures discussed herewith so far. We toughly value clear setting form lower segment of the economy as it serves building block in an entire sectorial formation. The study is aimed to examine the effect of households’ stock of social capital on their economic performance. Besides we also aim to examine the relative importance of general and special social capital stock in explaining the trend of economic growth using household data from Jimma zone. The study covers cross-section of about 383 units for 2018/2019.

METHODOLOGY

Study Area

This household survey was confined to Jimma Zone of Oromia regional state, situated at the South-west of Ethiopia (Figure 1). According to Jimma zone Finance and Economic Development office report (2016), Jimma zone is composed of 21 sub-administrative units called woreda. According to the same report, the total population of Jimma zone was estimated over three million under 512,506 households. The climatic condition of Jimma zone is characterized by 15% highland, 67% midlands and about 18% lowlands (UNDP, 2000). Jimma is among the major coffee growing regions in the country. The area consistently receives good rain, ranging from 1200-2800 mm per annum. Therefore, this gives an opportunity to invest in agri-business and grantees the feasibility of these projects in particular. Besides, Jimma city serves as the largest market center in South-West Ethiopia. With easy access to transportation and other necessities, Jimma serves adequate market center for potential investments in and around its geographical boundary. It has latitude and longitude of about 7°40′N and 36°50′E, respectively. Jimma is characterized to exhibit tropical rainforest climate as of the Köppen climate classification. Temperature at Jimma ranges between 20°C and 25°C year-round range, on daily average (Lewis, 2014). Below is an official map of Jimma Zone (UNDP, 2000).

![FIGURE 1. AN OFFICIAL MAP OF JIMMA ZONE (UNDP, 2000)](image-url)
Data Type and Source
This study is based on both quantitative and qualitative data forms. The quantitative element was directly obtained from the respondents. Informations related to infrastructure access, distances from nearest market place, total annual production, total annual expenditure and income levels, household sizes and other relevant measures were collected directly from the included respondents. Supportive data such as Zonal marketing report, statistical reports from Central Statistical Agency, Rural Development Bureau of Jimma zone and various woreda level offices were also used as our secondary sources. Qualitative information related to the specific channels through which social capitals affect households’ economic outcomes were also obtained from the individual respondents included for the purpose of our survey.

Since the study was based on information both from primary and secondary sources, family-centric informations such as; social capital endowment and economic position, were collected through structured questionnaires. Data on previous commodity price as well as production levels were used from the records of Jimma zone trade and Industry Department to supplement our secondary information needs.

Study Population and the Sample
The rural households from Jimma zone are the target groups with the present survey. We move through various stages to arrive at the final unit included in the analysis. We have started by arbitrarily selecting four out of twenty one woredas making up Jimma zone. These include; Limmu-Kossa, Sokoru, Manna and Gommaworedas. Again, each selected woreda was sub-divided in various sub-administrative units called Kebele, and required number of kebele was randomly selected from each of the four woredas. Finally, proportional size across sample woreda was arbitrarily taken from respective kebele administrations.

According to official reports of respective woreda administration, household size across each of the four woreda were reported to be, 25,305 for Limu Kossa; 30,115 for Sokoru; 18,201 for Manna; and 21,449 for Gomma. The total sum of households in four districts is equals 95,070. Due to (Noelet al., 2012), the sample size selected from the four districts was 383 (equation 1).

\[
\frac{n}{N} \geq \frac{N}{1 + (N-1)(\frac{d}{z})^2}
\] .............................. (1)

Where N stands for total population; n represents sample size; (d =0.05) is margin of error; and (z=1.96) for a 95 per cent confidence interval.

The required sample size was proportionally distributed across each sample woreda as follows;

- Limu Kossa: \( \frac{383}{95,070} \times 25,305 = 102 \) households
- Sokoru: \( \frac{383}{95,070} \times 30,115 = 121 \) households
- Manna: \( \frac{383}{95,070} \times 18,201 = 74 \) households
- Gomma: \( \frac{383}{95,070} \times 21,449 = 86 \) households
Data Analysis

We rely both on descriptive and inferential statistical tools to analyze the data. Descriptive tools indicating themaximum and minimum values as well as mean and standard deviation of various measures were used. Besides, binomial logit regression was employed to estimate the impact of a household total stock of social capital on economic performance. Hence, the latter replicates the inferential aspect.


In this paper we contend that, social network is a vital element of social capital. Social capital may be defined as stock of social relationships that an individual person possesses (Ferragina, 2010; Putnam, 1995). Individuals can form relations with others in different ways. Hence, social relations can be grouped as general and special social. General relations are typically congenital from genetic factors. An important feature of such relationships is that they are established and are already there for you whether you like it or not. As far as they are established genetically, you will have no contribution in your share the general aspect of social capital. Besides, at their work place individuals could interact with peer as well as their bosses by default. In this case, general relation is already around whether/or not a worker likes. Relations established externally without individual effort are features of general social capital (Wickramasuriya et al, 2011; Akcomak, 2008; Knowles, 2005).

Sometimes any one member of a given family may give special trust to another member, which he cannot share with the rest in the family. A member with such special respect and trust from a group is said to have special capital on relative terms. Similarly, from a class of say, sixty or more, one or two may have close relation and communication to the school director, not all. Obviously, those with better relation to their director will have better access to some key school facilities or, even personal contribution from the director. Turning back to our point, special aspect of social capital begins from recognition of such seemingly small but significant interpersonal relations between individuals. A boss my not treat all staff members equally; he may share only functional roles with all, but core issues can be communicated to some with relatively better personal quality. The special aspect of social relation gives better trust and certainty than the already established general relation. To clarify the difference among the two, general social relations are large in size but shallow in depth; while the special aspects are in converse (Wickramasuriya et al, 2011; Shufang and Yan, 2010; Thokozani, 2012).

To simplify things, the general relation based on blood is limited to family membership; i.e., sharing the same mother or father or either. Based on (Saidov, 2017; Miller, 2003; OECD, 2009; Shufang and Yan, 2010; Ponthieux, 2004) general aspect of social capital has been measured by; (a) Number of people a household possess relation to; and (b) Engagement in civic organization. We used ‘Idir’ as a local formation to proxy the civic engagement in the local sense. Following (UN, 2005; Thokozani, 2012; Wickramasuriya et al, 2011) we capture the special component of social capital guided by whether one is quite sure that he or she has, at least, one person in life that he or she will not leave him or her during difficulties in all life events; good or bad, long lived or short lived. In the same passion above,
sum total of the two components form a household’s total social capital stock (TSSC). Stated differently, the social capital can be split into two as general and social capital stock; for an ith individual respondent.

With this conceptualization, a households’ total stock of social capital (TSSC) is, thus, the sum total of general social capital (GSC) and Special Social Capital (SSC). It follows that;

\[ TSSC_i = GSC_i + SSC_i \]  \hspace{1cm} (2)

2. Measuring Household Economic Wellbeing

With regard to household economic wellness, we used the individual household annual income and expenditure accounting based on (Sharma et al., 2012). The wellbeing analysis refers to whether a household is able to cover personal livelihood without being indebted each year. In this conception, if a household has never borrowed for family consumption during the survey period then that household is economically safe. To determine each household current net income, we have systematically collected information on past year production, sale, consumption, and expenditure (in terms of ETH) item by item throughout the year. A difficulty with our approach was that some products were nonmarketable; and some others were not sold in the same year, difficulty to have their money value. A notable difficulty arises from the fact that, considerable proportion of rural production is allotted for family consumption and that; a proportion couldn’t sell in the same year they are produced. Efforts were made to come-up with handling mechanisms: productions at hand are monetized at respective current market price. Henceforward, an aggregate of annual spending has been withheld from gross income obtained in that year. The difference account for the periodic economic position of each household surveyed. Then, total expenditure was deducted total income earned during the year.

\[ NET_i = INC_i - EXP_i \]  \hspace{1cm} (3)

Where, \( NET \) stands for net income of ith household; \( INC_i \) and \( EXP_i \), respectively, are the annual income and expenditures of household i. Since the household being economically secure is the response variable in logit specification, we need to define it clearly. Whether or not a household is economically safe is categorical in two layers. The response variable assumes 1 value if a household under consideration is economically secure; and 0 otherwise. Finally, the response variable can be defined as follows;

\[ E.SEC_i = \begin{cases} 1, \text{ if a household is secure} \\ 0, \text{ otherwise} \end{cases} \]  \hspace{1cm} (4)

3. Binary Logistic Regression

We model the actual effect of social capital on household economy via the approach for binary logistic analysis. Households’ economy is a response variable in the logit specification. If the computed net income for a household is positive, corresponding income dummy takes on a value 1; and zero otherwise. The specification is based on the notion that, with positive net income a household is able to cover current consumption expenditure without being indebted.

An outcome variable \( y \) follows a Bernoulli distribution taking 1 (occurrence) with probability of success \( P \), or 0 with probability of failure \( 1-P \) in the logistic regression
framework. In this paper logit was specified letting \(y_{(x)}\) be a dichotomous with classes 1 (economic security) and 0 (no economic security).

Due to (Maddala, 1998; Wooldridge, 2013), the vectors of predictors (\(x\)) and outcome variable (\(y\)) can be developed as shown below:

SEM is also a stronger hybrid technique in terms of confirmatory aspects of factor analysis, path analysis and regression or considered as having interaction modeling, nonlinearity, correlated independent variables (Narimawati & Sarwono, 2017).

\[
X = \begin{bmatrix}
1 & x_{11} & \cdots & x_{1k} \\
1 & x_{21} & \cdots & x_{2k} \\
\vdots & \vdots & \ddots & \vdots \\
1 & x_{nk} & \cdots & x_{nk}
\end{bmatrix}
\]

\[
Y = \begin{bmatrix}
y_1 \\
y_2 \\
\vdots \\
y_n
\end{bmatrix}
\]

\[
\beta = \begin{bmatrix}
\beta_0 \\
\beta_1 \\
\vdots \\
\beta_k
\end{bmatrix}
\sim (k + 1) \times 1
\]

Where \(X\) is vector of regressors; \(Y\) is a column vector of an outcome variable; \(\beta\) represents the vector of parameters. Following Gujarati (2004), the probability that a household is economically secure for known values of each regressor considered is given by;

\[
P_i = \frac{1}{1 + e^{-Z_i}} \quad \text{................................. (5)}
\]

Where \(Z_i = \beta_0 + \beta_i X_i\); \(P_i\) is a conditional probability that the \(i^{th}\) household is economically secure for the known values of each regressor in the model; \(\beta_0\) is a constant; \(\beta_i\) are slope coefficients; and, \(e\) is base of natural logarithm. For mathematical treatments for negative exponents, (equation 5) can be re-expressed as follows;

\[
P_i = \frac{e^{Z_i}}{1 + e^{Z_i}} \quad \text{................................. (6)}
\]

On the other hand, a probability that a household is economically insecure equals the value leftover of equation (6) is deducted from unity;

\[
1 - P_i = \frac{1}{1 + e^{Z_i}} \quad \text{................................. (7)}
\]

The odds ratio in favor of a household to be economically secure is given by the ratio of the two probabilities.

\[
P_i / (1 - P_i) = \frac{e^{Z_i}}{1 + e^{Z_i}} = e^{Z_i} \quad \text{................................. (8)}
\]

The natural log of equation (8) gives the below final expression;

\[
L = \ln \left( \frac{P_i}{1 - P_i} \right) = \ln (e^{Z_i}) \Rightarrow \beta_i + \beta_i X_i + U_i \quad \text{................................. (9)}
\]

Where \(L = \text{Log of odds ratio and } u_i\) is the error term. Based on equation (9) we model the following final relation between the dependent and independent variables of this study.

\[
L = \beta_0 + \beta_1 \text{TSSC}_i + \beta_2 \text{IDIR}_i + \beta_3 \text{EDUC}_i + \beta_4 \text{AGE}_i + u_i \quad \text{................................. (10)}
\]

Where TSSC is Total Stock of Social Capital of household; EDUC is Education level of the respondent \(i\) measured in years of schooling; AGE is the age in years of the \(i^{th}\) household head; and IDIR is a dummy taking a value of 1 if a respondent is member of its community ‘Idir’; and 0, if not a member.
4. Diagnostic Test

Once the model is fitted, we would be interested to know how effective the model is in explaining the outcome variable.

a. Pearson Chi-Square Test

The Pearson $\chi^2$ statistic is based on observed ($O$) and expected ($e$) observations.

$$\chi^2 = \sum_{i=1}^{n} \frac{(O_i - e_i)^2}{e_i} = \sum_{i=1}^{n} \frac{(y_i - n_i \hat{P}_i)^2}{n_i \hat{P}_i (1 - \hat{P}_i)} \text{ .................................................. (11)}$$

Where $y_i$ is the observed value of $Y_i$; $\hat{P}_i$ is fitted $Y_i$ for a given value of $x_i$; $n_i$ is the number of observations. Large Pearson chi-square for a given regressor indicates that there is strong association between each of independent and the dependent variables keeping the effect of the other factors constant. Here we test the null hypothesis of $H_0$: there is no association between variables against its alternative $H_1$: there is association between the dependent and particular independent variable describing the outcome variable.

b. Wald Test

A Wald test is used to test the statistical significance of each coefficient ($\beta$) in the model. If the Wald test is significant for a particular explanatory variable, then we would conclude that the parameter associated with the variable is not zero so that the variable should be included in the model; otherwise it should be omitted from the model (Agresti, 1996). We test the null hypothesis of $H_0: \beta_j = 0$ against the alternative $H_1: \beta_j \neq 0$ in the Wald test. The Wald test statistic ($Z$) for this hypothesis is

$$Z^2 = \frac{\hat{\beta}_j^2}{\text{var}(\hat{\beta}_j)} \sim \chi^2 (1) \text{ .................................................. (12)}$$

Where $\hat{\beta}_j$ is the estimated regression coefficient and $\text{var}(\hat{\beta}_j)$ is the variance of $\hat{\beta}_j$.

c. The Likelihood Ratio Test

The likelihood ratio test statistic ($G^2$) is the test statistic commonly used for assessing the overall fit of the logistic regression model. The LL ratio is computed based on $-2LL$ (2 times log likelihood). The likelihood ratio statistic is obtained by subtracting the two times log likelihood ($-2LL$) for the full model from the log likelihood for the intercept only model. This log likelihood-ratio test uses the ratio of the maximized value of the likelihood function for the intercept only model $L_0$ over the maximized value of the likelihood function for the full model $L_1$. The likelihood test statistic is given by:

$$G^2 = -2 \log \left( \frac{L_0}{L_1} \right) = -2[\log(L_0) - \log(L_1)] = -2[LL_0 - (LL_1)] \text{ .................................................. (13)}$$

Where, $LL_0$ represents the log likelihood value of the model which has the intercept term only and $LL_1$ is the log likelihood value of the full model. The likelihood ratio statistic
has a chi-square distribution and it tests the null hypothesis that all logistic regression coefficients except the constant are zero. The degrees of freedom are obtained by differencing the number of parameters in both models. It is compared with chi-square value at the difference between degree of freedom of both models. The p-value less than 5 percent level of significance lead to the rejection of the null hypothesis thereby implying that all the predictor effects are zero. When the likelihood test is significant, at least one of the predictors is significantly related to the response variable (Hosmer and Lemeshow, 2000).

d. Hosmer-Lemeshow Test

The Hosmer-Lemeshow test is used to check the overall model fit. In this approach, data are divided into ten groups. From each group, the observed and expected numbers of events are computed. Then, the Hosmer-Lemeshow test statistic is given by:

$$
\hat{c} = \frac{g}{V_j} \sum_{j} (O_j - E_j)^2 
$$

Where $E_j = np_j$; $g$ is the number of groups; $O_j$ is observed number of events in the $j^{th}$ group; $E_j$ is expected number of events in the $j^{th}$ group and $V_j$ is a variance correction factor for the $j^{th}$ group. If the observed number of events differs from what is expected by the model, the statistic $\hat{c}$ will be large and there will be evidence against the null hypothesis that the model is adequate to fit the data. This statistic has an approximate chi-square distribution with $(g-2)$ degrees of freedom (Agresti, 1996).

e. Other Tests

In logistic regression, outliers are observations whose values deviate from the expected range and produce extremely large residuals. These outliers can unduly influence the results of the analysis and lead to incorrect inferences. An observation is said to be influential if removing the observation substantially changes the estimate of coefficients. Influence can be thought of as the product of leverage and outliers. An observation with an extreme value on a predictor variable is called a point with high leverage. Leverage is a measure of how far an independent variable deviates from its mean. In fact, the leverage indicates the geometric extremeness of an observation in the multi-dimensional covariate space. These leverage points can have an unusually large effect on the estimate of logistic regression coefficients (Cook, 1998). Using the following rules it is possible to identify if an observation is outlier or influential:

- **Residuals**: Standardized, Standard, deviance and Pearson residuals are obtained using different software. Observations with values larger than three in absolute values are considered as outliers (Agresti, 2007).
- **DFBETAS**: Measure of how much an observation has affected the estimate of a regression coefficient (there is one DFBETA for each regression coefficient, including the intercept).
Cook’s D: It measures an average effect on the group of regression coefficients, as well as the group of fitted values. In logistic regression, a case is identified as influential if its Cook’s distance is greater than one (Hosmer & Lemeshow, 2000).

RESULT AND DISCUSSIONS

Descriptive Information

1. Household Socio-Demographic Characteristics

A sample of 383 households from Jimma zone was arbitrarily taken to address the objectives of this study. We have collected information on various issues; like, the demographic characteristics, socio-economic and political conditions describing our target group. The family size and age distribution across our sample was presented in table (1).

<table>
<thead>
<tr>
<th>Variable</th>
<th>Obs.</th>
<th>Mean</th>
<th>Std. Dev.</th>
<th>Min</th>
<th>Max</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age</td>
<td>383</td>
<td>35.6</td>
<td>6.9196</td>
<td>19</td>
<td>56</td>
</tr>
<tr>
<td>Family Size</td>
<td>383</td>
<td>4</td>
<td>0.8514</td>
<td>1</td>
<td>8</td>
</tr>
</tbody>
</table>

Source: (Own Survey, 2018)

Evident from table (1) is high age variability with standard deviation of about 6.92. The range of the age variable equals 37 with minimum and maximum 19 and 56 years, respectively. Therefore, we can say that the study has considered the sections of the society at various stages; hence, that could further guarantee the reliability of our analysis results thereby simultaneously avoiding biases arising from concentration of the sample on either extreme of the age group. Besides, the mean age is about 35 years and 6 months showing that majority of the study population were young.

The mean size of family was estimated around four with standard deviation of 0.8514 thereby signifying less variability in size distribution. This result may also imply that, any social, economic and political differences across households resulting from differences in family size cannot be an issue with us. As far as the family size distribution was symmetric, the multi-dimensional welfare differences among individual households because of differences in size might not be important. The mean family size is four; while the minimum and maximum sizes are 1 and 8, respectively.

Table (2) reports the religion and gender backgrounds of the households included in the current analysis. The data on the sex and religion variables belongs to the household heads. It should be noted that, the religion of household head belongs also to every member of the family.

The framework of the corn farmers’ performance behavior model, especially farmers’ cooperatives, as a liaison between agro-food supply chains of small-scale producers and sustainable institutions for gaining more added values, is determined by the farmers’ characteristics. The participants were mostly 40 to 45 years old; their ethnic is java. They had been married for more than 20 until 25 years, they had 4 family members, their education background is 9 years, and most of them had 2-10 year farming experience.
### Table 2. Religion and Sex Distribution of the Sample Respondents

<table>
<thead>
<tr>
<th>Sex category</th>
<th>Frequency</th>
<th>Percentage</th>
<th>Cum. Pert.</th>
</tr>
</thead>
<tbody>
<tr>
<td>Male</td>
<td>103</td>
<td>26.89</td>
<td>26.89</td>
</tr>
<tr>
<td>Female</td>
<td>280</td>
<td>73.11</td>
<td>100</td>
</tr>
<tr>
<td>Gross</td>
<td>383</td>
<td>100</td>
<td></td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th></th>
<th></th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td>Muslim</td>
<td>127</td>
<td>33.16</td>
<td>33.16</td>
</tr>
<tr>
<td>Orthodox</td>
<td>80</td>
<td>20.89</td>
<td>54.05</td>
</tr>
<tr>
<td>Protestant</td>
<td>122</td>
<td>31.85</td>
<td>85.90</td>
</tr>
<tr>
<td>Wake feta</td>
<td>54</td>
<td>14.10</td>
<td>100</td>
</tr>
<tr>
<td>Gross</td>
<td>383</td>
<td>100</td>
<td></td>
</tr>
</tbody>
</table>

*Source: (Own Survey, 2018)*

As it can be observed from table (2), out of 383 total respondents, only 26.89 per cent (103) of the total response rate are male; while the remaining 73.11 per cent (280 out of 383) belongs to female counterparts. That means about 73.11 per cent of the household was female-headed. Hence, most of the household decision in the study area was made by females.

Table (2) also reveals that majority of the respondents were followers of Islam accounting for about 33.16 percent of the total response rate. The second dominant religion was Protestantism representing about 31.85 percent of the total response rate. 122 out of 383 total respondents are Protestants. Besides, about 80 (20.89%) individuals were reported to follow Orthodox Christianity; while, the remaining 14.10 percent of the total response rate belong to Wake feta. There is no such large group variability regarding religion distribution between Islam and Protestantism as they both account for almost equal proportion of the total response rate. On the one hand, the proportion Wake feta fellowship was comparatively trifling. Therefore; Islam, Protestantism and Orthodox Christianity are widely practiced in the study area. We realize religious heterogeneity and coexistence of diversified views was also acknowledged from Jimma zone. The proportion of various religions from Jimma zone can also be shown using simple pie-chart as displayed hereunder.

![Pie Chart](chart.png)

*Source: (Own Survey, 2018)*

**Figure 2. Frequency Distribution of Religion among Sample Respondents**

Figure (1) is a reflection of table (2), which again indicates the dominancy of Islam followed by Protestantism, Orthodox Christianity and Wake feta respectively.
2. Education Level and Ethnic Composition

Ethnic composition in Jimma zone was found to be heterogeneous thereby holding dominantly of Amhara, Dawro, Oromo and Yem ethnic groups among others. However, the area is most dominantly inhibited by the Oromo people accounting for more than half estimated at about 56.13 percent of the total response rate. It means that, out of 383 sample respondents 215 individuals belong to Oromo ethnic group indicating the dominancy of the indigenous group. Amhara was the next dominant group thereby accounting for about 24.02 percent of the total response rate. Of the total of 383 sample respondents, 92 belong to the Amhara ethnic group. The third largest share goes to the Yem people accounting for about 16.71 percent of the total response; while the remaining 3.39 percent of the total response rate goes to Dawro people in the general ethnic composition of the study area. More than 80 percent of the total respondents belong to only two ethnic groups of which about more than 70 percent is explained by the Oromo people. Therefore, we found the dominancy of one ethnic group in the study area revealing large variability in the distribution of ethnic group.

<table>
<thead>
<tr>
<th>TABLE 3. EDUCATION AND ETHNIC COMPOSITION</th>
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<tbody>
<tr>
<td>Indicators</td>
</tr>
<tr>
<td>Ethnic Composition</td>
</tr>
<tr>
<td>Amhara</td>
</tr>
<tr>
<td>Dawro</td>
</tr>
<tr>
<td>Oromo</td>
</tr>
<tr>
<td>Yem</td>
</tr>
<tr>
<td>Total</td>
</tr>
</tbody>
</table>

| Education Level | Frequency | Percentage | Cum. Pert. |
| Diploma | 39 | 10.18 | 10.18 |
| Secondary Education | 104 | 27.15 | 37.33 |
| Primary Education | 42 | 10.97 | 48.30 |
| Illiterate | 198 | 51.70 | 100 |
| Total | 383 | 100 | |

Source: (Own Survey, 2018)

We can also demonstrate the ethnicity and education level frequency of households using diagrams. Figure (2) below indicates the frequency distribution of ethnic composition as discussed above.
More than half of the total response rate is illiterate indicating the proportion of individuals who didn’t attend at least basic education. There is evidence that out of 383 total respondents, 198 are illiterate exceeding half of the total sample by more than 6 individuals. A proportion of individual respondents who pursued secondary education totals 104; which is about 27 percent of the total response rate; while 42 out 383 total respondents attended primary education. Alternatively, about 11 percent of the total response rate has pursued primary education. Furthermore, a proportion of individual respondents out of the total response rate to have graduated in diploma were about 10.18 percent. Taking head count of individual respondents, it means only 39 out of 383 total respondents were diploma graduates. Considering the large size of cross sectional observation with this study this size is very small.

Generally, the frequency distribution analysis of education indicates that majority, more than half of the total response rate was illiterate. This indicates, most of household consumption, investment and marketing decisions are inefficient; given the fact that, education would help in intensifying the effectiveness and quality of human activity. Besides, it is this higher illiteracy rate that is responsible to explain the prevailing multidimensional poverty in the area. More than half of the total response rate didn’t have access to at least basic education could imply that most household activities are based on non-technical traditional practices. Inefficient decisions reduce labor and other resource productivity as well as the overall product level. Poor education may be one of the basic factors to explain their weak economic, political and other dimensional positions. Using figures, we can equivalently analyze the frequency distribution of education among the sample respondents as follows;

![Figure 4. Education Pursued by Households](source: Own Survey, 2018)

We observe from figure (3) that, the highest room for the illiterate group being followed by those who attended secondary education. Both groups together form about 78.85 percent of the total response rate of which about 65.56 percent are still illiterate. Those who pursued tertiary education account for a very small proportion of the total response rate. Overall, this particular analysis on the frequency distribution of education among the sample household reveals that, majority of the individuals in the study area is illiterate. We could hence infer that, this high illiteracy rate may be a factor that could explain multidimensional poverty and inadequate welfare across households from Jimma zone.

**Inferential Analysis**

Prior to the estimation of binary logistic model, we have addressed the necessary validity, consistency and reliability requirements issues with logistic regression. In none of the
specific cases was the necessity test failed. We present the individual test result in sections below.

Let’s first see the model fitness condition. As discussed under methodology section, the goodness-of-fit of the logistic specification can be verified by the reported log likelihood ratio together with the regression coefficients, the score condition and the Wald test. Each test of the overall model strength has proved valid; and hence the inference based on logistic estimates.

<table>
<thead>
<tr>
<th>Test List</th>
<th>Chi2</th>
<th>Pr &gt; Chi2</th>
</tr>
</thead>
<tbody>
<tr>
<td>LL Ratio</td>
<td>57.4782</td>
<td>&lt;.0001</td>
</tr>
<tr>
<td>Score</td>
<td>59.3547</td>
<td>&lt;.0001</td>
</tr>
<tr>
<td>Wald</td>
<td>52.6925</td>
<td>&lt;.0001</td>
</tr>
</tbody>
</table>

Source: Own Manipulation using STATA ver. 14 (2018)

We see that (table 4), the Likelihood ratio, Score and Wald test statistics are all significant with p-value less than 0.05. This indicates that the final model with explanatory variables included was more effective than a model without explanatory variables. Hence, there is significant relationship between the dependent and set of all independent variables. Moreover, we have examined the Hosmer-Lemeshow condition for the goodness-of-fit of our logistic specification, and confirmed the fulfillment. The Hosmer-Lemeshow test statistic was estimated at 12.65 (p=0.1244), which is strongly significant at 5 percent.

With regard to logistic diagnostic issue, we considered DFBETA for each regressor as well as the average term, standard residuals and influential observation issues, and all were tested robust. The Adequacy of fitted model was examined for possible presence and treatment of outliers and influential observation. See table (5) below for the test statistics.

<table>
<thead>
<tr>
<th>Test Type</th>
<th>Obs.</th>
<th>Minimum</th>
<th>Maximum</th>
</tr>
</thead>
<tbody>
<tr>
<td>Analog of Cook’s influence statistics</td>
<td>383</td>
<td>3.30e-08</td>
<td>0.0450509</td>
</tr>
<tr>
<td>Standard Residual</td>
<td>383</td>
<td>-2.110746</td>
<td>1.85213</td>
</tr>
<tr>
<td>DFBETA for TSSC</td>
<td>383</td>
<td>-0.3675358</td>
<td>0.1578752</td>
</tr>
<tr>
<td>DFBETA for EDUC</td>
<td>383</td>
<td>-0.2547097</td>
<td>0.3009825</td>
</tr>
<tr>
<td>DFBETA for AGE</td>
<td>383</td>
<td>-0.2025523</td>
<td>0.2216249</td>
</tr>
<tr>
<td>DFBETA for IDIR</td>
<td>383</td>
<td>-0.1789904</td>
<td>0.1874204</td>
</tr>
</tbody>
</table>

Source: Own computation using stata ver. 14 (2018)

The DFBETAs for parameters and Cook’s influence statistics are both less than unity. DFBETAs of less than unity indicating no specific impact of a given observation on the coefficient of a particular predictor variable; whereas Cook’s distance less than one reveals that an observation had no universal effect on vectors of logistic regression coefficients (\( \hat{\beta} \)). The normalized residuals were reported to lie within the interval of (-3 and 3), demonstrating the absence of an outlier in our logistic specification. Consequently, relevant statistical requirements have tested robustness with our logistic specification of this study.

1. **Effect of Social Capital Stock on Household Economy**

Here we aim to examine the impact of households’ aggregate social capital stock on their economic performance, measured by the method briefed so far. As it is quite clear, a
The principal regressor in our logistic specification is social capital stock. The ‘Idir’ dummy accounted by the household civic engagement was another aspect of social capital in the model. Besides, we also control for household heads’ number of years of schooling as well as age in years. Our Stata 14 display for logistic manipulation looks a copy below:

TABLE 6. SOCIAL CAPITAL AND HOUSEHOLD ECONOMY

<table>
<thead>
<tr>
<th>Regressors</th>
<th>Coeff.</th>
<th>Std. Error</th>
<th>Prob.</th>
</tr>
</thead>
<tbody>
<tr>
<td>TSC</td>
<td>0.311</td>
<td>0.050</td>
<td>0.000</td>
</tr>
<tr>
<td>IDIR</td>
<td>0.416</td>
<td>0.307</td>
<td>0.229</td>
</tr>
<tr>
<td>AGE</td>
<td>0.154</td>
<td>0.031</td>
<td>0.176</td>
</tr>
<tr>
<td>EDUC</td>
<td>0.150</td>
<td>0.041</td>
<td>0.000</td>
</tr>
</tbody>
</table>

Log LL Ratio: Chi² (4) = 87.55 [0.000] Pseudo R² = 0.2125

Source: Own Survey (2018)

For logit estimate reliability issues, the log likelihood ratio indicating overall relevance of the model is robust thereby ensuring reliability of inferences based on our estimates. Besides, regressors entered the model jointly account for about a quarter of variations in the response variable, signifying the need for consideration of other variables for future studies in related topics.

Evident from table (6) is the strongly significant and positive coefficient of household’s total stock of social capital on its economic status. Its estimated coefficient being 0.311 (0.000) signifies that, when an individual household’s stock of social capital increases by a unit, its odds in favor of economic wellbeing will be intensified by about 0.311 units. Besides, an improvement in the households’ total stock of social capital increases the likelihood that an individual household is economically secured. Moreover, the estimated weight of social capital is relatively large thereby further revealing its relevance in the wellbeing of the people. The economic aspect of social capital could be realized as a result of production efficiencies (Nie et al., 1996; Durlauf, 2002; Loury, 1977); decreased transaction costs (Greif, 1993; DiPasquale and Glaeser, 1999); better team work and effects of individual networking (Ball et al., 1998; Putnam, 1995(2000)). This finding is theoretically expected and empirically similar to Reimer (2014) in Canada; Reinsberg et al (2013) in Poland; and Terrence (2014) in United States. An Idir dummy was suggested to have insignificant effect on household economy. The reported differential effect (0.416) was suggested insignificant even at 10 per cent. Thus, there are no real economic differences across households whether or not they are members of their community self-help association called Idir. It means that the dummy for Idir has not been important in explaining the households’ economic position. This result may not be surprising from the economic perspective; given that, such associations, like Idir, are mainly targeted for social rather than economic matters. With Idir, individuals come together to share social emergencies and occasional burdens in life.

Education was another relevant consideration. We have estimated a positive (0.150) and statistically significant (p=0.000) impact of household’s education level on its economic position. The coefficient reveals that, a year extra schooling increases odds of household being economically secure by 0.15, which is statistically significant improvement from the computational point of view. Hence, household economic status has been found to vary...
positively with the level of education attended. The positive coefficient of education is obvious. Education affects life adequately; and in respect of economy, education is argued to enhance mechanized production techniques thereby developing knowledge and workers and improving managerial efficiency. All these allow for individual’s raising productivity and hence overall product level.

An age variable was also found irrelevant. The estimated coefficient 0.014(0.229) reveals that the effect of householder’s age is not significant. It indicates that, when the age of a household head increases by a year, the odds of that household being economically secured improves by about 0.014; yet insignificant improvement. The connotation with this particular case may be that, life experience doesn’t matter while economic performance is considered. In fact, time allows an individual to learn while doing thereby simultaneously providing opportunities to take advantages of advancing technology that could enhance easy life. In contrast, if you let your time pass for nothing eventually you will finally end up with added but lost time life. Most of the time, experiences help in business and popularly claimed to contribute positive asset arising from economies of scale due to large size.

In general, the total stock of social capital that a household possess contributes towards improving its economic endeavor. Furthermore, households’ status in education was found important in explaining their economic performance. That means, with higher level of education households’ perform well in economic regards. But, Idi dummy and the age considerations were irrelevant with reference to household economy.

2. General versus Special Social Capital: Analysis of Relative Economic Importance

We took step ahead towards analyzing the relative importance of general and special components of social capital. Here, we examine as to which of the two household economic condition was more responsive. According to social science theories, the special aspect matters more compared to the general aspect, in reference to their multidimensional role in human life. The premise follows that, whether an individual likes or not, general relations are there and no effort from the individual to exist. In contrast, special relations are usually created with large effort, hence interpreted meaningfully. As a result, special aspect of social capital would be tough in human relationships (Ashford, 2005). The later aspect of social capital is more likely intensifying household economic as well as other relevant events, compared to the earlier. In this study, we expect better role of the special aspect of household social capital with reference to Jimma zone. To determine the relative importance, we run logistic regression in two regressors, and the result is displayed below:

| Regressors | Coef. | Std. Error | P>|Z| |
|------------|-------|------------|----------|
| GSC        | 0.2796505 | 0.2060365 | 0.175    |
| SSC        | 0.3455757 | 0.0567074 | 0.000    |
| Cons.      | -1.156174(0.000) | | |
| LR Chi²(2) = 69.12 [0.000] | Pseudo R² = 0.1678 | | |

Note: A value in [ ] is the probability of the log likelihood ratio.
Source: (Own Survey, 2018)
The validity of two regressor model can be ensured by its significant probability of chi square (table 7). Moreover, the reported pseudo-R2 value of about 16.78 percent is fairly good for such large cross sectional observation in only two regressors (Wooldridge, 2013; Poulton and Kent, 2008).

Figures in the table describe everything in advance! The special aspect of household social capital is more likely to significantly contribute towards household economic wellbeing (table 7). There is evidence for significant and positive effect of special social capital on household economic security. Our regression result reveals that, when an individual household’s special capital stock improves by a unit measure; the odds of economic security will be intensified by about 0.34; whereby by granting the likelihood of household economic safety. This improvement in the household’s economic position is highly meaningful as the estimated coefficient is strong. Besides, sign of the estimated coefficient was right; and empirically similar to (Ashford, 2005; Reimer, 2002; Reinsberg et al, 2005). The assertion is that, it is the special component of social capital that contributes more to the prosperity of the people. To the other side, households’ general social capital has been found not important in their economic considerations. Its estimated coefficient of 0.27(0.175) clearly reveals that the general aspect of household social capital has nothing to do with safeguarding their wellbeing. General relationships do not guarantee real interpersonal linkages, exchange of relevant ideas and meaningful interpretations of those relations; and that might explain why they often cease to maintain their individual as well as group welfare (Uzawa, 2005; Saidov, 2018; Dildar, 2010).

We see that general social capital has nothing to do, unless appropriated meaningfully. Obvious is a positive but insignificant impact from the households’ economic wellbeing; whereby the positive coefficient indicating if wise approaches were there for appropriation, it would have had desirable economic signal. In contrast, the role of special social capital in advancing household economy is strongly significant in the expected direction. The implication is straightforward; relations having no meaningful interpretations could end up non-sense.

CONCLUSIONS

This study established relationships between social capital and economic performance, using household survey from Jimma zone. A prime goal was estimating the impact of social capital on household economy. There is evidence that, an economic effect of social capital on household economic performance was positive and strongly significant. Separate logit regression yielded insignificant coefficient for general social capital, demonstrating irrelevant economic value from the general aspect of societal networks in the study area. But, the same regression indicated robust and positive sign for special social capital. Results tell us that, the economic relevance of special social capital was found much higher than the general stock. The implication follows that, though social capital plays obvious role for multifaceted social development, it could have nothing to do unless the share of special aspect is intensified. General social capital, which can be established exogenously without ones
desire, is not such influential. An important experience now calls for schemes to intensify growth of special relations even from general ones. Besides, interventions to encourage developments in social capital across households in rural Jimma could support the national program of ensuring food security and maximizing rural welfare.

REFERENCES


